

A Century of Gaps: The Global View

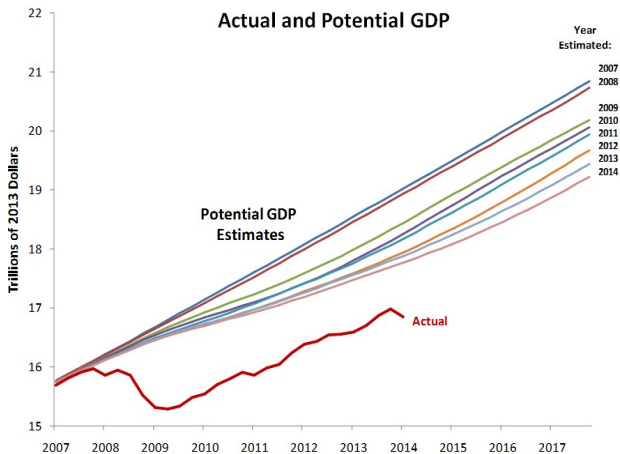
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September 27th 2018

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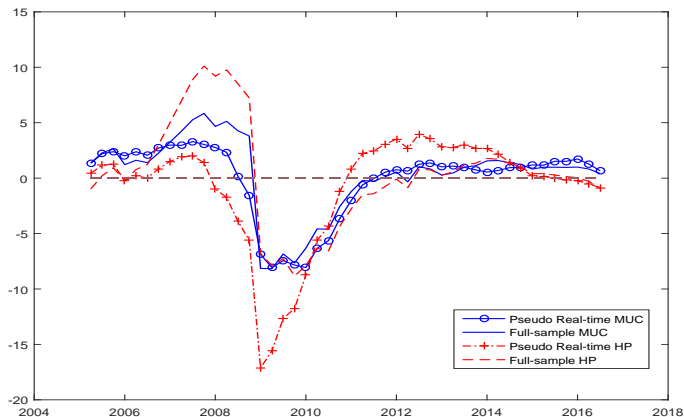
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The Motivation



- Propose new methods for the estimation of output gaps: **A Century of Gaps (2018)**, **Unemployment or Credit: Which One Holds the Potential? Results for a Small-Open Economy with a Low Degree of Financialization** (*forthcoming* Economic Systems)
- Assess the dynamics of country-specific gaps and their (statistical) drivers over more than 100 years: **Mind the Global Gap (2018)**

Real-time Output Gap: Lithuania



Notes: The pseudo real-time output gap is obtained by recursively estimating the models, MUC *versus* HP, using 1998:Q1-2005:Q1 as the initial sample and adding one by one observation from 2005:Q2-2016:Q3.

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 - 3 Multivariate unobserved component (MUC) model: Kuttner (1994), Blagrove et al. (2015), Alich (2015) and Melolinna and Toth (2016), Borio, Disyatat and Juselius (2014, 2016) - **BDJ**

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- A parallel stream of literature initiated by Schularick and Taylor (2012) provides substantial evidence of the role credit plays in driving boom and bust episodes (across time and space) - credit is a primary source of macroeconomic shocks not only a channel of shock propagation (usual implicit assumption in many financial-accelerator models)

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 - Borio et al. (2016, 2014) and Melolinna and Toth (2016) have argued that the inclusion of financial variables improve the real-time performance of measuring output gap.
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 - We explore the role of various financial factors in the estimation of output gap and potential output for the US, investigating the changing role of real credit growth over more than 100 years.
- ❸ Expand the original BDJ framework to allow financial variables to have a **time-varying impact** in the estimation of potential output and the associated gap. JST (2016) Macrofinancial History and the "New Business Cycle Facts" point to two distinct finance eras:
 - Pre-1940 features a stable relationship between money and credit, business driven credit creation process and a modestly varying credit-to-gdp ratio/
 - Post-1940 gives way to the "financial hockey stick": increasing aggregate leverage ratios with most credit going to mortgages.

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- ➌ The sensitivity of output gap to real credit growth is time-varying; the parameter is increasing since the beginning of 1990s
- ➍ The finance-neutral potential output has been constant at 2% since the beginning of the 1980s. Ignoring the financial variables severely distorts its level post-2007 crisis - Financial Drag Hypothesis seems more likely than the Secular Stagnation Hypothesis.

Let \bar{y}_t denote the unobserved trend component of log real GDP and \hat{y}_t its cycle (or output gap). Following Harvey and Todd (1983), Harvey (1985) and Harvey et al. (2007), the decomposition of y_t is written as follows:

$$y_t = \bar{y}_t + \hat{y}_t \quad (1)$$

The trend component is modelled as a BM

$$\Delta \bar{y}_{t+1} = \Delta \bar{y}_t + \varepsilon_t^{\bar{y}}, \text{ where } \varepsilon_t^{\bar{y}} \sim N(0, \sigma_{\varepsilon^{\bar{y}}}^2) \quad (2)$$

The output gap will depend on its own lag as well as on the evolution of a financial variable, here denoted f_t

$$\hat{y}_t = \rho \hat{y}_{t-1} + \gamma f_t + \varepsilon_t^{\hat{y}}, \text{ where } \varepsilon_t^{\hat{y}} \sim N(0, \sigma_{\varepsilon^{\hat{y}}}^2) \quad (3)$$

The output gap will depend on its own lag as well as on the evolution of a financial variable, here denoted f_t but now γ will be time-varying

$$\hat{y}_t = \rho \hat{y}_{t-1} + \gamma_t f_t + \varepsilon_t^{\hat{y}}, \text{ where } \varepsilon_t^{\hat{y}} \sim N(0, \sigma_{\varepsilon^{\hat{y}}}^2) \quad (4)$$

with

$$\gamma_t = \gamma_{t-1} + \varepsilon_t^{\gamma}, \text{ where } \varepsilon_t^{\gamma} \sim N(0, \sigma_{\varepsilon^{\gamma}}^2) \quad (5)$$

- Both models are cast in the state-space form allowing us to estimate the unobserved factor via the Kalman filter (Harvey and Todd (1983), Harvey (1985))
- Specify the prior of the parameters following the literature
- Estimate the mode of the posterior distribution by maximizing the log posterior function, which combines the prior information on the parameters with the likelihood of the data evaluated by Kalman filter.
- Sample from the posterior with a Markov Chain Monte Carlo to obtain the posterior distribution.

One of the potential problems when using real credit growth as a proxy for the financial cycle is that the trend in f_t may pass onto output gap estimates. To avoid this issue, we demean f_t before the estimation using a 10 year moving average.

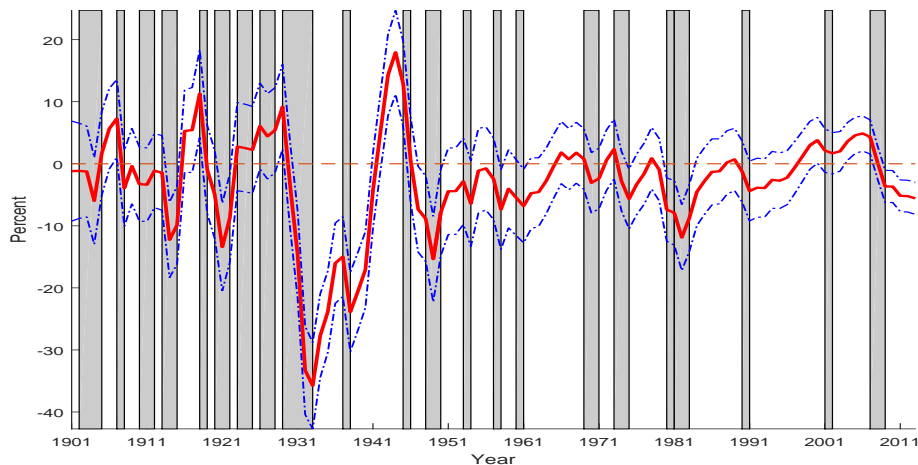
	Domain.	Prior distribution			Posterior Distribution		
		Distr.	Mean	St. Dev.	Mean	[5%, 95%]	PSRF
ρ	(0,1)	Beta	0.80	0.10	0.85	[0.76, 0.94]	1.00
$\sigma_{\varepsilon_{\bar{y}}}^2$	\mathbb{R}^+	Inv.Gam.2	1.00	1.00	0.58	[0.25, 1.17]	1.00
$\sigma_{\varepsilon_{\hat{y}}}^2$	\mathbb{R}^+	Inv.Gam.2	20.0	20.0	23.9	[18.9, 30.0]	1.00
γ	\mathbb{R}				0.24	[0.12, 0.37]	1.00

Notes: We do not restrict any prior belief on the parameter γ . The posterior distribution is obtained by the Metropolis-Hastings algorithm. PSRF- Potential Scale Reduction Factor. Real credit is obtained by deflating the total loans to non-financial private sector by CPI.

The TVP-BDJ model allows the information content that financial factors have for output gap to change over time by modelling the financial parameter γ_t as a driftless random walk process.

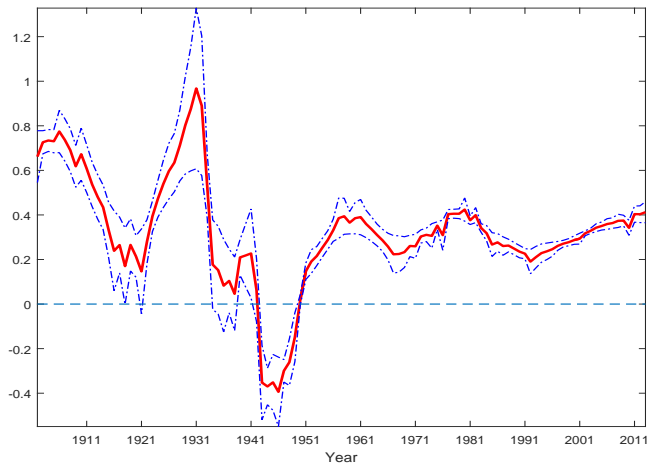
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$\sigma_{\varepsilon_{\tilde{y}}}^2$ \mathbb{R}^+	Inv.Gam.2	1.00	1.00	0.56	[0.25, 1.09]	1.00
$\sigma_{\varepsilon_{\hat{y}}}^2$ \mathbb{R}^+	Inv.Gam.2	20.0	20.0	15.8	[11.9, 20.4]	1.00
$\sigma_{\varepsilon_{\gamma}}^2$ \mathbb{R}^+	Inv.Gam.2	0.10	1.00	0.05	[0.02, 0.10]	1.00

Figure: Output Gap Estimates (± 2 SE)



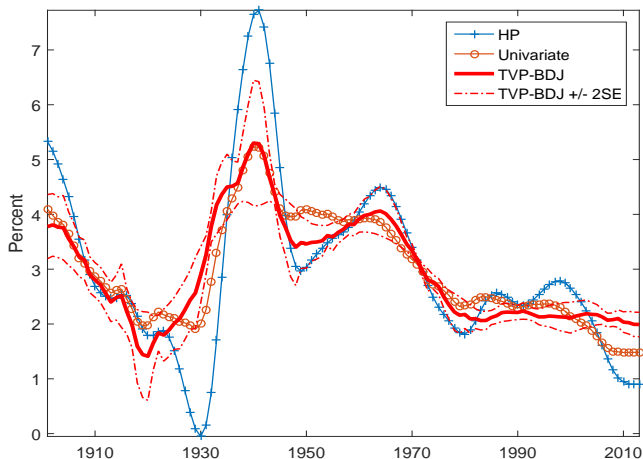
Note: Estimates of output gap, which is 100 times natural log deviation of output from its potential level using the *TVP-OG* model. Shaded areas: recessions (NBER)

Figure: Real Credit Growth Time-Varying Parameter (± 2 SE)



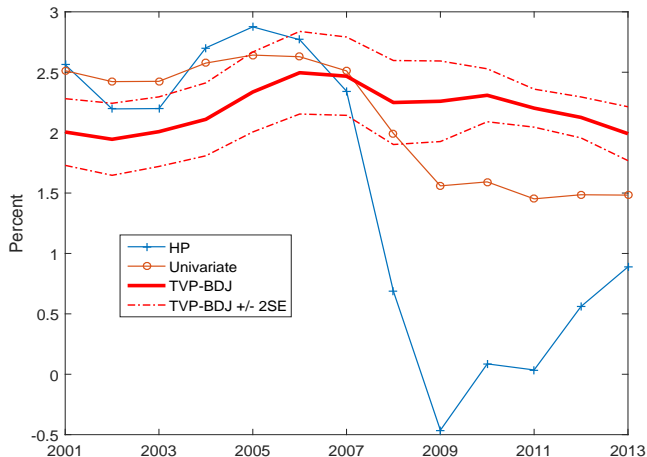
Note: This figure presents the time-varying influence of real credit growth on output gap γ_t .

A comparison of methodologies

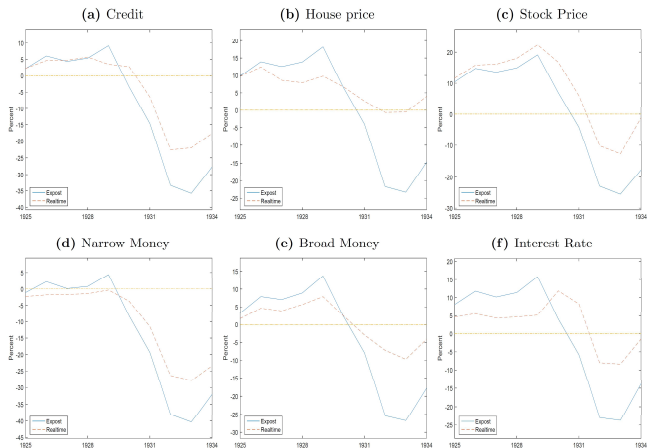


Notes: The figure shows the smoothed estimates of the potential growth of output ($\Delta \bar{y}_t$) from three different models: the HP filter, the univariate model, and the TVP-OG model.

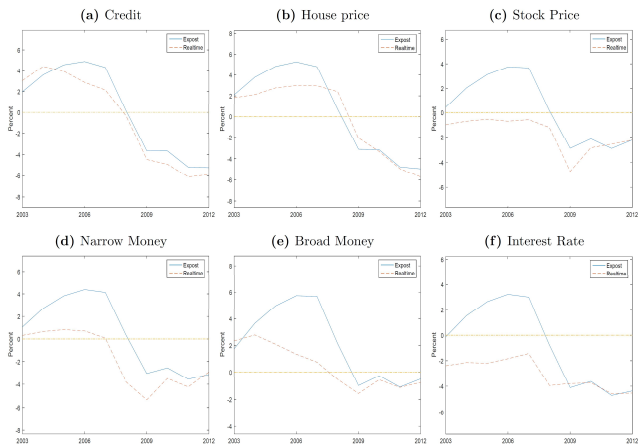
Potential Growth - Zooming In



Notes: The figure shows the filtered estimates of the potential growth of output ($\Delta \bar{y}_t$) from three different models: the HP, the univariate model, and the TVP-BDJ model. The filtered estimates are those estimated using the sample of data available up to the point of estimation. For the HP, it is the one-sided HP estimate with the smoothing parameter being equal to 100.



Note: The figure shows the real-time (dashed) and ex-post (solid) output gap estimates for the 1925-1934 period using the TVP-BDJ framework with different proxies for financial information.



Note: The figure shows the real-time (dashed) and ex-post (solid) output gap estimates for the 2003-2012 period using the TVP-BDJ framework with different proxies for financial information.

Mind the Global Gap!

In Constantinescu and Lastauskas (2018), we find an important role of cross-sectional averages of GDP, housing and credit as covariates explaining national house price dynamics over more than 120 years of economic history - with different impact over time and in the cross-section.

- We provide a century-long view of the contribution of global and national factors to the cyclical component of credit, economic activity, housing and equity markets for all the countries in the JST (2016) and Knoll (2017) dataset.
- We estimate a dynamic multivariate latent factor model similar to Mumtaz (2017) to uncover the increasing importance of global factors at *medium-term* rather than short-term frequencies.
- Given the increasing importance of global factors, our work provides *evidence of the need for global coordination* in both macro-stabilization and macro-prudential policies.
- A high degree of heterogeneity is identified across countries and time in responses to global components.

- Schularick and Taylor (2012, 2017): *credit is a primary source of macroeconomic shocks not only a channel of shock propagation.*
- The empirical relevance of global factors in the determination of business cycle commonalities across countries has been highlighted in several papers employing dynamic multivariate latent factor models such as Kose, Otrok, and Whiteman (2003) and Del Negro and Otrok (2008).
- More recently, the increasing interdependence of real and financial cycles as well as its heterogeneous nature across both time and countries feature prominently in studies of global economic converge such as Kose, Otrok, and Prasad (2012), Claessens, Kose, and Terrones (2012), Mumtaz (2017).

- Are there common factors potentially at work in both real and financial variables? If yes, policy may benefit from more consideration of joint dynamics of macroeconomic stabilization **and** macro-prudential policies.
- Do effects from common components manifest themselves in the short-term or rather over longer time-frames? Borio (2012) "Characterising the financial cycle: dont lose sight of the medium term!" - if important, effects of QE might surprise us with some unexpected latent effects in the years to come

- Real output to estimate business cycles (BC)
- Real total loans to non-financial private sector for credit cycles (CC)
- Real house prices for house price cycles (HC)
- Stock price for stock price cycles (SC).²

The data is annual spanning the period 1870 to 2013 from the Macrophistory database. We consider 16 countries as shown in Table 1.³ Four cycles are estimated for each country, but Italy and Spain⁴

Table: List of countries

Australia	Belgium	Canada	Denmark
Finland	France	Germany	Italy
Japan	Netherland	Norway	Spain
Sweden	Switzerland	UK	US

²The real term is obtain by deflating the nominal term by CPI.

³Portugal is not included because both stock price and house price are not available over a large portion of the estimation sample.

⁴In the latter, house prices are only available from 1970 so the series are excluded from these two countries.

- Comin and Gertler (2006) document the existence of the medium-term components in GDP which tend to be swept into the trend by conventional filters: rather than working with the filter at frequencies between 2 and 8 years, authors argue in favor of 2 to 20 years.
- Recent emerging studies on the interaction between business cycles and financial cycles document that the financial cycle is much longer than the traditional business cycle and business cycle recessions are much deeper when they coincide with the contraction phase of the financial cycle (Drehmann et al. (2012)).
- To cast light on the importance of the two different frequencies, we estimate the weight of global and national factors (along with idiosyncratic shocks) at both short and medium-frequencies.

We assume that each cycle X_{it} is decomposed according to a dynamic factor model as follows:

$$X_{it} = B_{it}^C F_t^C + B_{it}^W F_t^W + u_{it}. \quad (6)$$

Each factor k of F_t is described by:⁵

$$F_{kt} = \sum_{i=1}^p b_{ki} F_{kt-i} + e_{kt}, \quad e_{kt} \sim N(0, 1). \quad (7)$$

The idiosyncratic component u_{it} is assumed to follow an $AR(q)$ process:

$$u_{it} = \sum_{j=1}^q d_{ij} u_{it-j} + e_{it}, \quad e_{it} \sim N(0, \sigma_i). \quad (8)$$

Factor loadings on the country and world factors, $B_{it} = [B_{it}^C, B_{it}^W]$, are allowed to be time-varying, based on a random walk process:

$$B_{it} = B_{it-1} + \tau_t, \quad \tau_t \sim N(0, Q_i). \quad (9)$$

⁵The standard deviation of e_{kt} is set to one, which is a standard normalization assumption because of the scaling issue.

The variance of X_{ct}^i can be decomposed as follows:

$$\text{var}(X_{ct}^i) = \sum_{k=1}^{K^C} (B_{k,ct}^i)^2 \text{var}(F_{k,t}^i) + \sum_{j=1}^{K^W} (B_{j,ct}^{i,W})^2 \text{var}(F_{j,t}^W) + \text{var}(u_{ct}^i). \quad (10)$$

Based on Equation 10, we can evaluate the contribution of each country factor $F_{k,t}^i$ to the variance of X_{ct}^i by

$$CV_{k,ct}^i = \frac{(B_{k,ct}^i)^2 \text{var}(F_{k,t}^i)}{\sum_{k=1}^{K^C} (B_{k,ct}^i)^2 \text{var}(F_{k,t}^i) + \sum_{j=1}^{K^W} (B_{j,ct}^{i,W})^2 \text{var}(F_{j,t}^W) + \text{var}(u_{ct}^i)}, \quad (11)$$

and the total contribution of country factors to the variance of X_{ct}^i is:

$$CV_{ct}^i = \sum_{k=1}^{K^C} CV_{k,ct}^i. \quad (12)$$

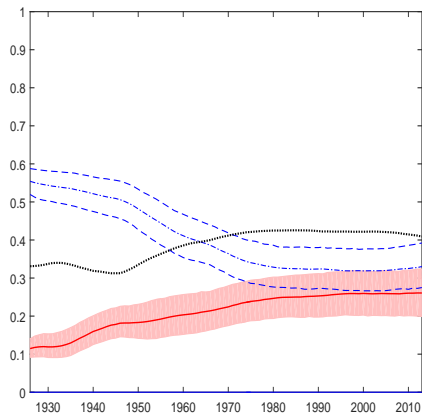
Similarly, the contribution of each global factor $F_{j,t}^W$ to the variance of $X_{c,t}^i$ is calculated by

$$GV_{k,c,t}^i = \frac{(B_{j,c,t}^{i,W})^2 \text{var}(F_{j,t}^W)}{\sum_{k=1}^{K^C} (B_{k,c,t}^i)^2 \text{var}(F_{k,t}^i) + \sum_{j=1}^{K^W} (B_{j,c,t}^{i,W})^2 \text{var}(F_{j,t}^W) + \text{var}(u_{c,t}^i)}. \quad (13)$$

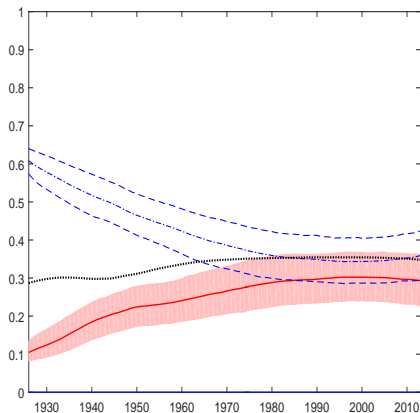
and the total contribution of global factors to the variance of $X_{c,t}^i$ is:

$$GV_{c,t}^i = \sum_{j=1}^{K^W} GV_{j,c,t}^i. \quad (14)$$

Short-term Average Across Countries: BC and CC



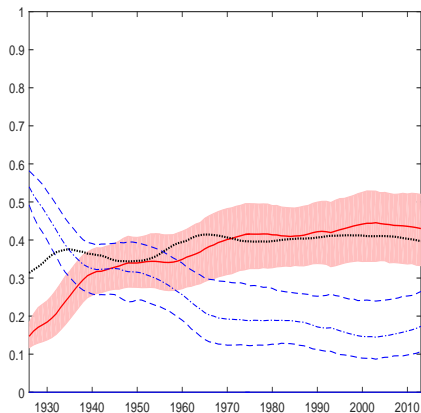
(a) Business



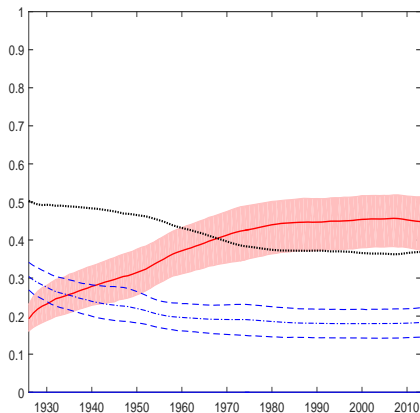
(b) Credit

Note: The contribution to the variance of each series of global factors (red line with 90% error bands - red shaded area), country factors (blue dashdot line with 90% error bands - blue dashed lines) and other factors (black dotted line).

Short-term Average Across Countries: HPC and SPC



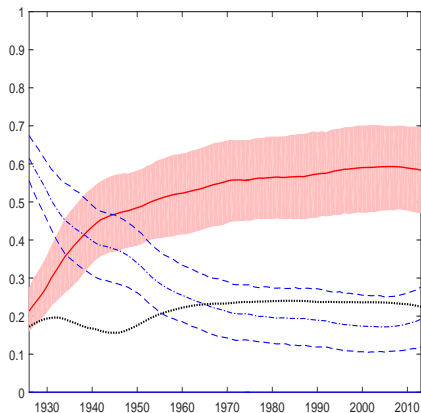
(c) House Price



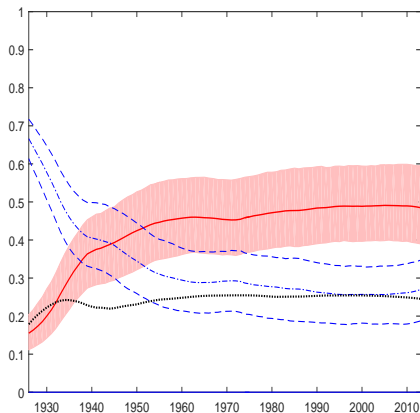
(d) Stock Price

Note: The contribution to the variance of each series of global factors (red line with 90% error bands - red shaded area), country factors (blue dashdot line with 90% error bands - blue dashed lines) and other factors (black dotted line).

Medium-term Average Across Countries: BC and CC



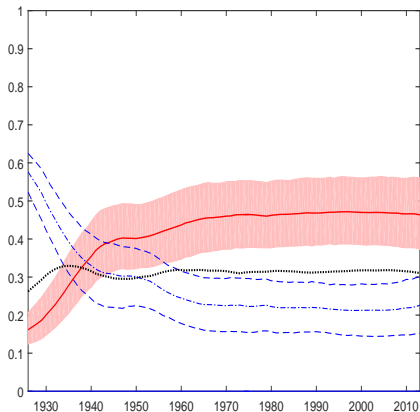
(e) Business



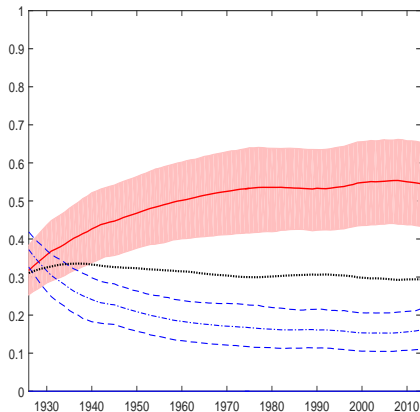
(f) Credit

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Medium-term Average Across Countries: HPC and SPC



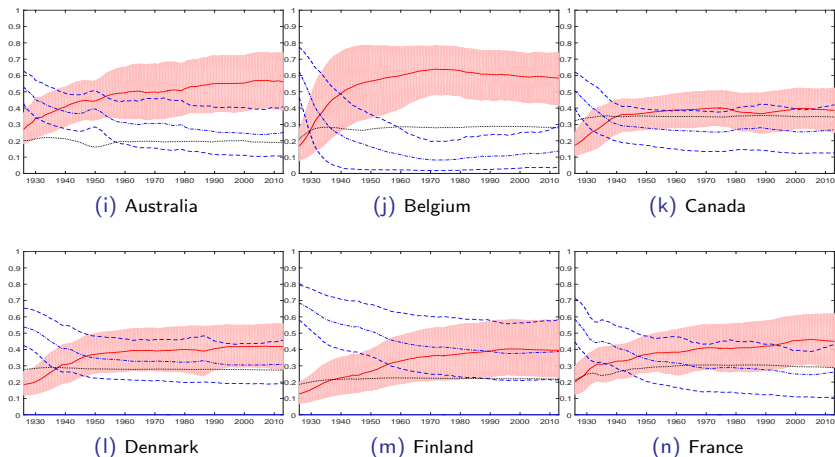
(g) House price



(h) Stock price

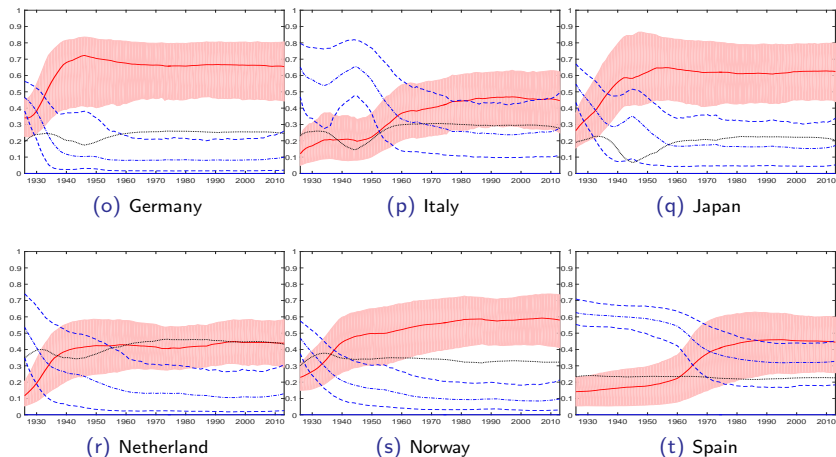
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Country-averages: medium term cycle



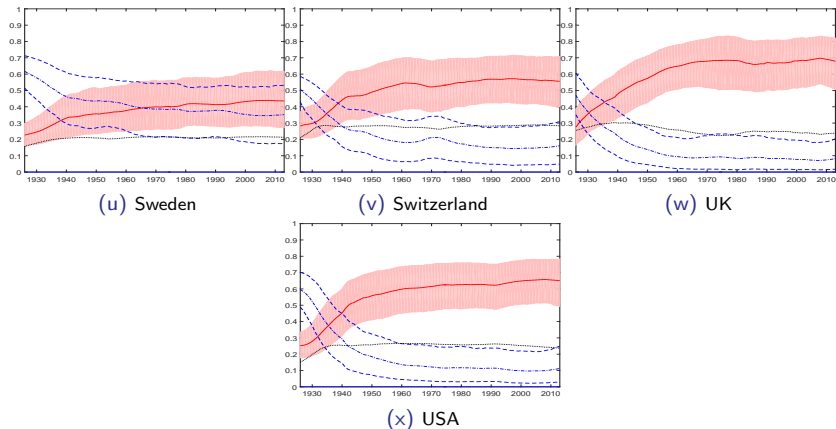
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Country-averages: medium term cycle



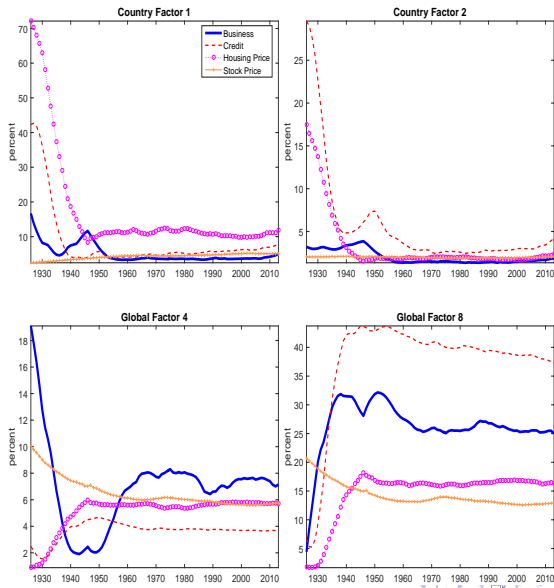
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Real-Financial Interaction in Contribution to Variance: Germany



Real-Financial Interaction in Contribution to Variance: US

